# Estimating Detection Probabilities and Site Occupancy Rates of Three Anuran Species Using Call Surveys in Haenam Gun, Korea 

Sung, Ha-Cheol*, Su-Kyung Kim, Seok-Wan Cheong, Shi-Ryong Park, Dong-Chan Roh ${ }^{1}$, Kyung Whan Baek², Jung-Hyun Lee ${ }^{3}$ and DaeSik Park ${ }^{3}$<br>Department of Biology Education, Korea National University of Education, Cheongwon, Chungbuk 363-791, Korea<br>${ }^{1}$ Jeonbuk Science High School, Iksan, Junbuk 570-911, Korea<br>${ }^{2}$ Haenam High School, Haenam, Junnam 536-803, Korea<br>${ }^{3}$ Department of Science Education, Kangwon National University, Chuncheon, Kangwon 200-701, Korea


#### Abstract

We investigated the distribution of three anuran species, Three-striped pond frogs (Rana nigromaculata), Bullfrogs (Rana catesbeiana), and Narrow-mouthed toads (Kaloula borealis), in an administrative district, Haenam Gun, Junnam Province, Korea using volunteer call surveys. Twenty-eight volunteer call surveyors were assigned to each $2 \times 2 \mathrm{~km}^{2}$ survey plot. Call surveys on whether the species are present or not were conducted for 5 minutes between 30 minutes after sunset and the midnight on rice fields and ponds from 10 April to 28 August in 2005. Depending on species, call surveys were carried out at seven to 28 plots with average 8.4 to 10.7 visits per the plot. We calculated the detection probabilities and occupancy rates of the three species using four models with three covariates: temperature, humidity, and the amount of water at the habitat. The model average detection probabilities of three anuran species of $R$. nigromaculata, R. catesbeiana, and $K$. borealis were $0.53,0.74$, and 0.41 respectively, and the site occupancy rates of them were $0.93,0.94$, and 0.86 respectively. Our results indicate that R. nigromaculata, R. catesbeiana, and K. borealis are common in Haenam Gun.


Key words: Anuran species, Call survey, Kaloula borealis, Rana catesbeiana, Rana nigromaculata, Site occupancy rate

## INTRODUCTION

The declines of amphibian populations are worldwide for the past several decades (Cooke 1972, Wake 1991, Alford and Richards 1999, Houlahan et al. 2000). Several factors for the amphibian declines have been suggested such as habitat destruction and modification, invasive species, chemical contaminants, and global climate change (Blaustein and Wake 1990, Storfer 2003). To elucidate the status of amphibian populations and the causes of such declines, extensive monitoring efforts have been made (Bell et al. 2004, Pellet and Schmidt 2005).

Call survey monitoring of amphibian populations based on volunteers is mainly used to cover a wide range of areas with low costs and with ease (McDiarmid and Donnelly 1994, Campbell et al. 2002, MacKenzie 2005, Pellet and Schmidt 2005). From the call surveys, site occupancy probabilities as well as metapopulation dynamics with colonization and extinction rates and turnover rates can be examined. Recently, MacKenzie et al. (2002) developed a method to allow unbiased estimation of detection probability and
site occupancy rate of the species using call survey data. For this method, volunteers need to perform multi-visits at a site and need to collect several covariances such as temperature and humidity information.
The goals of this study were to investigate regional distribution of three anuran species, Three-striped pond frogs (Rana nigromaculata), Bullfrogs (Rana catesbeiana), and Narrow-mouthed toads (Kaloula borealis), within an administrative district, Haenam Gun, Junnam Province, Korea. In this paper we estimated site occupancy rates and detection probabilities of the three species based on repeated surveys of a site within a breeding period and calculated the least number of visits necessary to confirm the absence of a species at a site. This is the first documentation of large-scale monitoring programs in Korea a step toward developing statistical models to reflect the current status of the three species as well as to predict environmental changes with long-term surveys.

## MATERIALS AND METHODS

## Study Area and Survey Methods

[^0]Call surveys were conducted by volunteers of Haenam high school students. In order to help volunteers discriminate among anuran calls, we provided a published booklet, which contains morphological, physiological, ecological, and acoustic information on amphibian species and methods of anuran call monitoring. A schoolteacher trained students to discriminate $R$. nigromaculata, $R$. catesbeiana, and $K$. borealis's calls. Total 31 volunteers participated in this study and we analyzed the data of 28 volunteers due to incomplete data records of three students.

Volunteers surveyed anuran calls once per week for R. nigromaculata and R. catesbeiana from 10 April to 28 August in 2005 and for $K$. borealis from 24 April to 7 August in 2005 in Haenam Gun, Junnam Province, Korea. Thus weather conditions were differed between the two breeding periods of three species: mean temperature $\pm$ standard deviation were $22.4 \pm 4.2$ (range, $9-!$ ) for $R$. nigromaculata and R. catesbeiana while $23.0 \pm 3.0$ (range, 1723); mean humidity $\pm$ standard deviation were $66.5 \pm 14.1$ (range, $32-$ i) for R. nigromaculata and R. catesbeiana while $69.7 \pm 11.5$ (range, $47-$ ). We divided the Haenam Gun area into $2 \times 2 \mathrm{~km}^{2}$ plots: total 150 survey plots of which more than $50 \%$ areas were land were produced. Out of total 150 plots, 28 plots (19\%) for $R$. nigromaculata, 18 plots (12\%) for R. catesbeiana, and 7 plots (5\%) for $K$. borealis were covered (Fig. 1). Surveyors heard anuran calls for 5 min between 30 minutes after sunset and the midnight on rice fields or on ponds and recorded the following information: fivelevel calling index ( 0 - no calls detected; 1-calls detected from one male; 2 - calls detected from two to five males; 3-calls detected from six to ten males; 4 - calls detected from more than ten males), weather, temperature, humidity, and three-level water index on the field ( 0 - less amount of water than usual; 1-usual amount of water; 2-more amount of water than usual). In this study to estimate site occupancy rates, we reassigned five-level calling index into either presence (1) or absence (0).

## Data Analysis

We calculated, first, naïve site occupancy rates (the number of detected plots divided by total study plots per species) and, second, weekly call detectability across study plots to overlook species detectability in Haenam Gun areas before entering data into a monitoring analysis program. In order to obtain the average plots of detected species per week, we calculated the sum of weekly rates of the detected over the total plots and divided the sum by the number of survey weeks, for which we used data from 17 April to 31 July for R. nigromaculata and R. catesbeiana and from 1 May to 31 July for $K$. borealis due to not enough numbers of surveyors. Then, we estimated detection probabilities and site occupancy rates with program PRESENCE (available from http://www.proteus.co.nz),


Fig. 1. Map of anuran call survey locations in Haenam Gun, Junnam Province, Korea. The study area was divided into 150 plots with a $2 \times 2 \mathrm{~km}^{2}$ plot, where 28 plots for $R$. nigromaculata, 18 plots for R. catesbeiana, and seven plots for K. borealis were surveyed by volunteered observers.
with which we consulted the data analysis modeled by MacKenzie et al. (2002). This program uses AIC (Akaike's Information Criterion) to select the most appropriate covariance model explaining the survey data; models in a candidate set were ranked by $\triangle$ AIC (the difference between the model with the lowest AIC and the given model); model weights ( $w$; relative likelihood of each model), indicating the degree of relative support of a model, were calculated; and then, we obtained model average estimates of detection probability based on the model weights (Burnham and Anderson 2002). Since this model assumes that sites are "closed" within a breeding season, we removed the data of the first and the last week surveys so that we confirmed the existence of the species within the survey plot during the survey periods.
We developed four models using three covariates to predict detection probability for each species: we assumed detection probability was constant ( $p$ ), or detection probability was associated with temperature, humidity, or three-level water index. To reveal how many visits are necessary to confirm the absence of a species at a site with $95 \%$ confidence, the following equation was applied (Pellet and Schmidt 2005):
$N_{\text {min }}=-\frac{\log (0,05)}{\log (1-p)}$
Where: $N_{\text {min }}$ is the number of visits; $p$ is estimated mean detection probability.

## RESULTS

## Rana nigromaculata

Surveyors visited their survey plots average 10.6 (range 4~16) times per plot including missing observations $45 \%$ ( 238 out of total 532 visits) for $R$. nigromaculata: the species was detected at least once at 26 out of 28 sites ( 0.93 ) which represents naïve site occupancy rates; and an average of 0.66 plots per week, where detectability was more than the average from 1 May to 25 June and peaked from 5 to 11 June (Fig. 2). Four models with three covariates showed similar weight ( $w$, Table 1). The overall estimated site occupancy rates ( $\hat{\mathscr{F}}$ ) were similar to naïve site occupancy rates $(0.93)$. As estimated mean detection probability of the models is 0.53 , minimum four visits are necessary to confirm that the species is absent.

## Rana catesbeiana

Surveyors visited their survey plots average 10.7 times (range $4 \sim 16$ ) including missing observations $44 \%$ ( 150 out of total 342 visits) for $R$. catesbeiana: the species was detected at least once at 17 of 18 sites ( 0.94 ); and an average of 0.65 plots per week, where detectability was more than the average from 22 May to 2 July except for one week of $12 \sim 18$ June and peaked from 5 to 11 June (Fig. 2). One model with Water as a covariate indicated that $R$. catesbeiana was more likely detected with increasing the amount of water (Fig. 3). Except that, all models showed similar estimates of site occupancy rate ( $\hat{\psi}$ ). As estimated mean detection probability of the models is 0.74 , minimum three visits are necessary to confirm that the species is absent.


Fig. 2. Weekly means ( $\pm \mathrm{SE}$ ) of the call detection probabilities of the three anuran species measured by volunteer based largescaled call surveys.

Table 1. Summary of models on parameter estimations of site occupancy rates and detection probabilities in the three species. $\Delta \mathrm{AIC}$ is the difference between the model with the lowest AIC and the given model; $w$ is the Akaike weight; $\hat{\psi}$ is the estimated proportion of sites occupied; $\operatorname{SE}(\hat{\psi})$ is the standard error of $\hat{\psi} ; \hat{p}$ is the estimated detection probability.

| Model | $\Delta \mathrm{AIC}$ | $w$ | $\hat{\mathrm{w}}$ | SE (产) | $\hat{p}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Rana nigromaculata |  |  |  |  |  |
| $\Psi() p.($ Temperature $)$ | 0.00 | 0.330 | 0.929 | 0.049 | 0.41 |
| $\Psi() p.($ Water $)$ | 0.12 | 0.310 | 0.929 | 0.049 | 0.43 |
| $\Psi() p.()$. | 0.46 | 0.262 | 0.929 | 0.049 | 0.71 |
| $\Psi() p.($ Humidity $)$ | 2.42 | 0.098 | 0.929 | 0.049 | 0.80 |
| Model averaged |  |  |  |  | 0.53 |
| Rana catesbeiana |  |  |  |  |  |
| $\Psi() p.($ Water $)$ | 0.00 | 0.853 | 0.945 | 0.054 | 0.79 |
| $\Psi() p.()$. | 4.98 | 0.071 | 0.944 | 0.054 | 0.64 |
| $\Psi() p.($ Temperature $)$ | 5.96 | 0.043 | 0.944 | 0.054 | 0.38 |
| $\Psi() p.($ Humidity $)$ | 6.54 | 0.032 | 0.944 | 0.054 | 0.39 |
| Model averaged |  |  |  |  | 0.74 |
| Kaloula borealis |  |  |  |  |  |
| $\Psi() p.($ Humidity $)$ | 0.00 | 0.357 | 0.857 | 0.133 | 0.30 |
| $\Psi() p.()$. | 0.10 | 0.340 | 0.858 | 0.132 | 0.48 |
| $\Psi() p.($ Temperature $)$ | 1.40 | 0.177 | 0.857 | 0.132 | 0.32 |
| $\Psi($ Water) | 2.06 | 0.127 | 0.858 | 0.132 | 0.66 |
| Model averaged |  |  |  |  | 0.41 |



Fig. 3. Relative frequency of absence and presence of R. catesbeiana for the three-level water index in Haenam areas over the study period.

## Kaloula borealis

Surveyors visited their survey plots average 8.4 times (range 5~ 14) including missing observations $38 \%$ ( 37 out of total 98 visits) for $K$. borealis: the species was detected at least once at 6 of 7 sites (0.86); and an average of 0.39 plots per week, where detectability was more than the average for two weeks from 1 May and for four weeks from 10 July and peaked for two weeks from 10 July (Fig. 2). All models showed similar estimates of site occupancy rate $\left({ }^{\hat{\psi}}\right)$. As estimated mean detection probability is 0.41 , minimum six visits are necessary to confirm that the species is absent.

## DISCUSSION

This study showed that detection probabilities varied among species but the site occupancy rates ( $\hat{\vec{F}}$ ) were constantly high among the four models in each species indicating that $R$. nigromaculata, R. catesbiana, and K. borealis are common in Haenam Gun. In addition, naïve site occupancy rates were also high for the three species. The difference between site occupancy rates and naïve site occupancy rates was zero or less than within the range of the standard error of each model, so that we suppose all plots we surveyed were found where the species existed. However, as the estimates of site occupancy rate and detection probability with presence/absence surveys are reasonable only for large-scale surveys (Yoccoz et al. 2001, Bailey et al. 2004), relatively limited survey areas for $K$. borealis, seven out of total 150 effective survey plots (5\%), may lead to overestimate those values for this species. For $K$. borealis, we may need further large-scale surveys in the future.

The results of this study may implicate that anuran call surveys of the three species using the models of MacKenzie et al. (2002) are useful in Haenam Gun area in two ways: First, as shown in the simulation results by MacKenzie et al. (2002), when surveyors visited more than five times at a survey plot with greater than 0.3 of detection probability, site occupancy rates appeared to be unbiased. Since our call surveys were performed averages of $8.4 \sim 10.7$ times per plot with larger than 0.41 detection probabilities, the estimates of the site occupancy rates must be accurate.

Second, averaged detection probabilities varied among the three species: 0.74 for $R$. catesbeiana, 0.53 for $R$. nigromaculata, and 0.41 for $K$. borealis. The various species detectability over the study plots may be correlated with overall calling activity; the weekly calling activity of $R$. nigromaculata and $R$. catesbeiana was relatively higher over the survey periods than that of $K$. borealis. The peaks of calling activity of both $R$. nigromaculata and $R$. catesbeiana appeared similarly in around 5 June but $R$. nigromaculata
was active early and prolonged till the peak while $R$. catesbeiana steadily increased the calling activity till a calling peak. K. borealis showed a prominent later calling peak around 10 to 17 July, the raining season in summer, than other species (Fig. 2). The limited seasonal calling activity of $K$. borealis may be responsible for relatively low detection probabilities of $K$. borealis than others. Nevertheless, call detectability of the species was satisfactorily high to estimate site occupancy rates. In addition, a model including Water as a covariate well predicted call detectability of $R$. catesbeiana while others appeared to be little important to predict that of $R$. nigromaculata and $K$. borealis. Previous researches on model selection of other anuran species showed temperature mainly played a role as a covariate in detecting the presence of the species (MacKenzie et al. 2002, Royle 2004, Pellet and Schmidt 2005). However, various environmental covariates also predicted the species detectability well in various ways among species and between years (Schmidt 2005): for example, 'day’ as a covariate best explained detection probability of Alytes obstetricans in 2001 while 'temperature' best explained in 2002.

To reduce time and cost and to increase the efficiency of call surveys, three recommendations may be useful for future surveys. First, since $R$. catesbeiana and $R$. nigromaculata can be confirmed to be absent at a site with minimum visits of $3 \sim 4$ times, four times visiting can give reasonable data. In case of $K$. borealis, since the species is seasonal specific, even though it needs to visit at least six times, we may conduct four times visiting only corresponding to the raining season. Second, to improve the detection probabilities, mating call playback methods may increase the detectability of three anuran species (Sung et al. 2005). Sung et al. (2005) showed a case study on $R$. nigromaculata that playbacks of the mating calls induced call utterances of silent resident males and higher calling activity from other males. Third, as Pellet and Schmidt (2005) pointed out, if the relationship between detection probability and a covariate exists, it may be able to determine the relevant temperature zone for effective detectability. For example, models including temperature as a covariate best predicted the variation of detection probability of Hyla arborea, where the species were detected well on warm nights that will be a good condition for surveys (Pellet and Schmidt 2005).

In this study, we obtained volunteer based large-scaled survey data on presence/absence of the three species at a site in Haenam Gun areas. The two estimates of detection probabilities and the site occupancy rates may provide more valuable information on the spatial distribution of the three species as well as on the temporal variations over several years in revealing the population status at local as well as regional levels, if we keep preceding the surveys and extending the survey areas.

## ACKNOWLEDGEMENT

Without the invaluable efforts by many volunteer students in Haenam high school, this study was not possible. We thank Ji Hye Yu, Seung Jae Lee, and Hyo Eun Jung for their help in the lab work. This study was supported by a grant (R01-2004-000-10450-0) from the Basic Research Program of the Korea Science \& Engineering Foundation

## LITERATURE CITED

Alford RA, Richards SJ. 1999. Global amphibian declines: a problem in applied ecology. Ann Rev Ecol Syst 30: 133-165.
Bailey LL, Simons TR, Pollock KH. 2004. Spatial and temporal variation in detection probability of Plethodon salamanders using the robust capturerecapture design. J Wildl Manage 68: 14-24.
Bell BD, Carver S, Mitchell NJ, Pledger S. 2004. The recent decline of a New Zealand endemic: how and why did populations of Archey's frog Leiopelma archeyi crash over 1996~2001? Biol Conserv 120: 189-199.
Blaustein AR, Wake DB. 1990. Declining amphibian populations: a global phenomenon? Trends Ecol Evol 5: 203-204.
Bridges AS, Dorcas ME. 2000. Temporal variation in anuran calling behavior: implications for surveys and monitoring programs. Copeia 2000: 587-592.
Burnham KP, Anderson DR. 2002. Model selection and multimodel inference: a practical information-theoretic approach, 2nd edition. Springer-Verlag, New York.
Campbell SP, Clark JA, Crampton L, Guerry AD, Hatch LR, Hosseini PR, Lawler JJ, O'Connor RJ. 2002. An assessment of monitoring efforts in endangered species recovery plans. Ecol Appl 12: 674681.

Cooke AS. 1972. Indications of recent change in status in the British Isles of the frog (Rana temporaria) and the toad (Bufo bufo). J Zool 167: 161-178.
Houlahan JE, Findlay CS, Schmidt BR, Meyer AH, Kuzmin SL. 2000.

Quantitative evidence for global amphibian population declines. Nature (London) 404: 752-755.
MacKenzie DI, Nichols JD, Lachman GB, Droege S, Royle JA, Langtimm CA. 2002. Estimating site occupancy rates when detection probabilities are less than one. Ecology 83: 2248-2255.
MacKenzie DI. 2005. Was it there? Dealing with imperfect detection for species presence/absence data. Aust New Zealand J Stat 47: 65-74.
McDiarmid RW, Donnelly MA. 1994. Group activities and field trips. In: Measuring and monitoring biological diversity-standard methods for amphibians (Heyer W, McDiarmid RW, Donnelly M, Hayek L, eds). Smithsonian Institute Press, Washington DC.
Pellet J, Schmidt BR. 2005. Monitoring distributions using call surveys: Estimating site occupancy, detection probabilities and inferring absence. Biol Conserv 123: 27-35.
Pollock KH, Nichols JD, Farnsworth G, Simons TR, Bailey L, Sauer JR. 2002. Large-scale wildlife monitoring studies: Statistical methods for design and analysis. Environmetrics 13: 1-15.
Richter SC, Young JE, Johnson GN, Seigel RA. 2003. Stochastic variation in reproductive success of a rare frog, Rana sevosa: implications for conservation and for monitoring amphibian populations. Biol Conserv 111: 171-177.
Royle JA. 2004. N-mixture models for estimating population size from spatially replicated counts. Biometrics 60:108-115.
Schmidt BR. 2005. Monitoring the distribution of pond-breeding amphibians when species are detected imperfectly. Aquatic Conserv 15: 681-692.
Storfer A. 2003. Amphibian declines: Future directions. Divers Distrib 9: 151-163.
Sung HC, Kim SK, Park SR, Park DS. 2005. Effectiveness of mating call playbacks in anuran call monitoring: a case study of Threestriped pond frogs (Rana nigromaculata). Integ Biosci 9: 199-203.
Wake DB. 1991. Declining amphibian populations. Science 253: 860
Yoccoz NG, Nichols JD, Boulinier T. 2001. Monitoring of biological diversity in space and time; concepts, methods and designs. Trends Ecol Evol 16: 446-453.
(Received May 4, 2006; Accepted May 26, 2006)


[^0]:    * Corresponding author, Phone: +82-43-230-3848, e-mail: shcol2002@hotmail.com

